Random Response Models or R-rated Questions of PG Audiences

General Introduction

Example: **(DON'T RESPOND)** How many of you are HIV positive?

Consider two methods of obtaining answers to this question:

- Heads up no privacy
- Heads on desks limited privacy

General areas where sensitive questions arise:

- Sexual behavior
- Drug or alcohol use
- Criminal actions
- Political or religous preferences

Probability

Population characteristics:

- Proportion with characteristic of interest is p. (*i.e.* 100p% HIV positive)
- Proportion without characteristic is 1 p.

Sample characteristics:

- Sample size is *N*.
- Response for the *ith* person may be represented by

| | | Response | |
|-----------------|--------|----------|------|
| Feature | Symbol | "Yes" | "No" |
| Random Variable | X_i | 1 | 0 |
| Probability | Pr | p | 1-p |

Simple Experiments

- Each X_i is called a *Bernoulli trial*.
- If we sum all of the X_i 's we get the total number of "Yes" responses. Denoted by $\sum_{i=1}^{n} X_i$.
- The average response, $\bar{X} = \sum_{i=1}^{n} X_i/N$, is the proportion of our sample which answered "Yes." It is also our best guess at the underlying proportion, p, of the population with the characteristic of interest.

Example: If p = 0.75 and N = 100 then

- Observing $\bar{x} = 0.72$ and $\bar{x} = 0.79$ wouldn't be too surprising.
- Seeing $\bar{x} = 0.2$ would be a bit out of the ordinary.

Example: If p = 0.75 and N = 2 then we can observe 0, 0.5 and 1 and nothing else.

- Repeated sampling gives us multiple estimates of p. The distributions of the sums from these samples are *Binomial*.
- The closeness of the estimates to the true proportion is measured by the *variance* of \bar{X} ,

$$\operatorname{Var}(\bar{X}) = \frac{p(1-p)}{N}$$

Example: For p = 0.75 and N = 100 we have variance

$$Var(\bar{X}) = \frac{0.75 \times 0.25}{100} \\ = 0.001875$$

Example: For p = 0.75 and N = 2 our variance is

$$Var(\bar{X}) = \frac{0.75 \times 0.25}{2}$$

= 0.09375

Direct Question Model

To find the proportion of HIV positive individuals we could ask each individual in our sample.

- Q_S = sensitive question (*e.g.* Are you HIV positive?)
- p = proportion of population with characteristic of interest
 - Estimated by \bar{X}
 - Variance is $Var(\bar{X}) = p(1-p)/N$

We write

$$\widehat{p}_{DQ} = \text{guess for } p$$

 $\equiv \overline{X}$

and

$$\operatorname{Var}\left(\widehat{p}_{DQ}\right) = \frac{p(1-p)}{N}$$
$$= \operatorname{Var}(\overline{X})$$

Can we trust the responses to Q_S ?

Randomized Response Model (Innocuous Question)

Notation:

- Greek letters for known probabilities: $\alpha,\,\pi$
- Latin letters for *unknown* probabilities: *r*, *p*, *a*
- Q_I = innocuous question (*e.g.* Is your favorite color blue?)
- α = proportion of population answering "Yes" to Q_I

Randomly assign a person to Q_S or Q_I . The interviewer is *blind* to the question being answered.

Example: Respondent flips a biased coin which the interviewer can't see. If a "Head" then respondent answers Q_S . If a "Tail" then Q_I is answered.

More notation:

- $\pi =$ probability respondent answers Q_S
- $1 \pi =$ probability respondent answers Q_I
- r_I = probability of a "Yes" response to the complete procedure
- \bar{X}_I = observed proportion of "Yes" responses to the randomized procedure

We know π and α and want to know p.

Our estimate of r_I is $\hat{r}_I = \bar{X}_I$.

Noting that

$$r_I = \pi p + (1 - \pi)\alpha$$

we have by substitution

$$\widehat{r_I} = \pi \widehat{p_I} + (1 - \pi)\alpha.$$

A little algebra brings us to

- our estimate, $\widehat{p_I} = [\widehat{r_I} (1 \pi)\alpha]/\pi$, of the proportion of the population with the characteristic of interest
- and our variance

$$\operatorname{Var}\left(\widehat{p_{I}}\right) = \frac{p(1-p)}{N} + \left(\frac{1-\pi}{\pi}\right)^{2} \left[\frac{\alpha(1-\alpha)}{N}\right] + \left(\frac{1-\pi}{\pi}\right) \left[\frac{\pi(1-\alpha) + \alpha(1-\pi)}{N}\right]$$

Example: Suppose we wish to check for the proportion of virgins within a population. We have p = 0.75 (unknown), $\alpha = 0.5$ (known), $\pi = 0.6$ (known) and N = 100 (known).

If the sample represents the population exactly, then

- 60 answer Q_S (45 are virgins)
- 40 answer Q_I (20 like blue)

Hence, $\bar{X}_I = \#$ "Yes" $= \frac{20+45}{100} = 0.65 = \hat{r}_I$

Thus, $\widehat{p_I} = \frac{0.65 - 0.4 \times 0.5}{0.6} = 0.45/0.6 = 3/4$

Audience Participation Example

Our questions:

- $Q_S =$ "Have you ever shoplifted?"
- $Q_I =$ "Was the card you cut a heart or a spade?"

Method: Respond to question

- Q_S if cut card is A–3 or 8–10
- Q_I if cut card is 4–7

Because the face cards were removed we have $\pi = 0.6$.

Since there are four suits in a deck we have $\alpha = 0.5$.

We compute

•
$$\bar{X}_I = \sum X_i / N =$$

•
$$\widehat{p_I} = \frac{\bar{X_I} - (1 - 0.6)0.5}{0.6}$$

Which leads to $\frac{\bar{X_I} - 0.2}{0.6} =$

Although we can't find the actual variance of this estimate we can get an estimate of the variance.

$$\operatorname{Var}(\widehat{p_{I}}) = \frac{p(1-p)}{N} + \left(\frac{1-0.6}{0.6}\right)^{2} \left[\frac{0.5(1-0.5)}{N}\right] + \left(\frac{1-0.6}{0.6}\right) \left[\frac{0.6(1-0.5)+0.5(1-0.6)}{N}\right]$$

If we cheat and substitute $\widehat{p_I}$ for p we get

$$\operatorname{Var}(\widehat{p_{I}}) \approx \frac{\widehat{p}(1-\widehat{p})}{N} + \frac{4}{9} \left[\frac{0.25}{N} \right] + \frac{2}{3} \left[\frac{0.5}{N} \right]$$

Associated Works

- Zellner, A. "An efficient method of estimating seemingly unrelated regressions and tests for aggregation bias," JASA June 1962, 348–368.
- Warner, S. "Randomized responses: a survey technique for eliminating evasive answer bias," JASA March 1965, 63–69.
- Abul-Ela, A., Greenberg, B., Horvitz, D., "A multiproportions randomized response model," JASA Sept. 1967, 990–1008.
- Gould, A., Shah, B., Abernathy, J., "Unrelated question randomized response techniques with two trials per respondent," Proceedings of the Social Statistics Section of the ASA 1969.
- Warner, S., "The linear randomized response model," JASA December 1971, 884–888.
- Cambell, C. and Joiner, B., "How to get the answer without being sure you've asked the question," The American Statistician, Dec. 1973, 229– 231.
- Dowling, T. and Shactman, R., "On the relative efficiency of randomized response models," JASA March 1975, 84–87.